

Model-Based Bioequivalence for sparse design pharmacokinetic studies: new statistical approach using Stan

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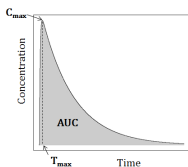
FDA: Andrew Babiskin, Sun Guoying, Stella Grosser, Liang Zhao, Lanyan Fang

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Bioequivalence (BE) studies

- **BE:** The difference in the pharmacokinetic (PK) of two formulations of a given drug does not exceed a predefined threshold (usually $\delta = \log(1.25)$)
- **Parallel studies** preferable for drugs with **long half-life:** $N/2$ subjects receive reference treatment (R), $N/2$ subjects receive test treatment (T)
- **Non-Compartmental Approach (NCA)** based BE:

- ▶ Individual NCA estimates of AUC_i and $Cmax_i$
- ▶ Let $\beta_{AUC} = \text{mean}(\log(AUC_{iT})) - \text{mean}(\log(AUC_{iR}))$
 $\beta_{Cmax} = \text{mean}(\log(Cmax_{iT})) - \text{mean}(\log(Cmax_{iR}))$



pros	cons
<ul style="list-style-type: none"> ★ Reproducible ★ Few assumptions 	<ul style="list-style-type: none"> ★ Require more than 10 samples per subject ★ Not appropriate for complex models

- **Two one-sided tests (TOST)**¹ at level $\alpha = 5\%$ on β_{AUC} and β_{Cmax} :
 Reject of H_0 (= BE significant) if
 - ▶ $(\beta - \delta) / SE(\beta) \leq -z_{1-\alpha}$ and $(\beta + \delta) / SE(\beta) \geq z_{1-\alpha}$, with $z_{1-\alpha}$: $(1 - \alpha)\%$ quantile of the normal distribution **or**
 - ▶ $CI(\beta)_{1-2\alpha\%} \in [-\delta; \delta]$

¹Schirmann, *J Pharmacokinet Pharmacodyn*, 1987.

Model based (MB) BE

Non linear mixed effect model (NLMEM):

y_{ij} : concentration for subject i at sampling time j

$$y_{ij} = f(t_{ij}, \phi_i) + g(t_{ij}, \phi_i) \epsilon_{ij}$$

$$\log(\phi_{il}) = \log(\lambda_l) + \beta_l Tr_i + \eta_{il} \text{ where}$$

- ▶ β_l : Test treatment effect on the log of a PK parameter $l = (1, \dots, p)$
- ▶ λ_l : l^{th} element of the vector of fixed effects
- ▶ Tr_i : vector of indicator for treatment group
- ▶ $\eta_{il} \sim \mathcal{N}(0, \omega_l)$: between subject random effect for parameter l
- ▶ $\epsilon_{ij} \sim \mathcal{N}(0, 1)$: residual error
- ▶ combined error model $g(t_{ij}, \phi_i) = a + b \times f(t_{ij}, \phi_i)$

Vector of population parameters: $\theta = (\lambda, \beta, \omega, a, b)$

- Estimation of θ using SAEM algorithm²
- Estimation of $SE(\theta)$ from observed Fisher information matrix Σ

²Kuhn et Lavielle. *Comput Stat Data Anal*, 2005.

MBBE

- AUC and Cmax: secondary parameters of the NLMEM
 - ▶ β_{AUC} and β_{Cmax} functions of vectors λ and β
 - ▶ $SE(\beta_{AUC})$ and $SE(\beta_{Cmax})$ obtained using the delta method³
- **TOST at level $\alpha = 5\%$ or 90%CI on β_{AUC} and β_{Cmax}**

pros	cons
★ Require few samples per subject	★ SE under estimated on sparse designs ⇒ Type I error inflation ⁴

- **Objectives: Development, evaluation and comparison of new model-based (MB) statistical approaches for sparse design BE studies**
 - ▶ Parametric random effect and residual bootstrap
 - ▶ Full distribution estimation using Stan

³Oelhert *The American Statistician*, 1992.

⁴Dubois et al. *Stat in Med*, 2011.

MBBE parametric random effect and residual bootstrap

Principle⁵

- 1 Estimation of θ and Σ with saemix
- 2 Drawing of $b = 1, \dots, B$ ($B=250$) matrices of random effects of size $N \times p$ from $\mathcal{N}(0, \hat{\Omega})$
- 3 Drawing of vector of residual errors of size $\sum_{i=1}^N n_i$ from $\mathcal{N}(0, 1)$
- 4 Simulation of the B vectors of responses
- 5 Fit the B new data sets with saemix to get the B estimates θ_b and β_b
- 6 Derive 90% CIs on β_{Cmax} and β_{AUC} from the 5th and 95th percentiles of the serie $\hat{\beta}_b$

⁵Thai et al, *J Pharmacokinet Pharmacodyn*, 2014.

MBBE full distribution using Stan

Principle⁶

- 1 Estimation of θ and η_i with saemix
- 2 Full distribution in Stan
 - ▶ Initialize HMC chain at estimates from step 1
 - ▶ Draw 1 chain of 1000 (including 100 burning) samples in *a posteriori* distributions of $\lambda, \beta, \omega, a, b$
 - ▶ out of the B=900 samples derive 90%CI on β_{Cmax} and β_{AUC}

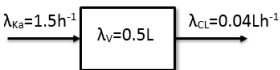
Stan model:

- Default distributions on fixed effects λ, β
- Non-informative priors on ω, a, b
 - $\text{omega_vec} \sim \text{cauchy}(0, 2.5);$
 - $a \sim \text{cauchy}(0, 2.5);$
 - $b \sim \text{cauchy}(0, 2.5);$

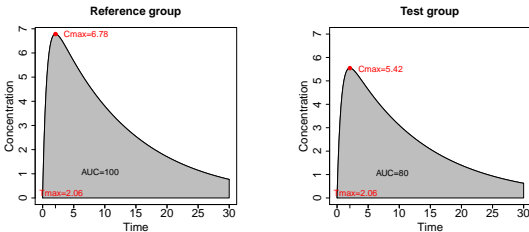
⁶Ueckert et al. *ACOP*, 2015.

PK model

- PK model of concentrations of the anti-asthmatic drug theophylline²
- One-compartment model with first-order absorption and first-order elimination



- Covariate effects in test treatment group: $\beta_V = \beta_{Cl} = \log(1.25)$, $\beta_{Ka} = 0$



- Residual variability: $a=0.1$ mg/L, $b=0.1\%$
- Random effects for BSV only:

ω_{ka} (%)	$\omega_{V/F}$ (%)	$\omega_{CL/F}$ (%)
22	11	22

²Dubois et al. *Stat in Med*, 2011.

Simulation scenarios

- 500 simulated data sets for each of the 4 scenarios

Design	N	Sampling times (hours)	Hypothesis	$\beta_{CL} = \beta_V$
Rich	40	n=10, $t = (0.25, 0.5, 1, 2, 3.5, 5, 7, 9, 12, 24)$	H_0	$\log(1.25)$
			H_1	$\log(1)$
Sparse	40	n=3, $t = (0.25, 3.35, 24)$	H_0	$\log(1.25)$
			H_1	$\log(1)$

Under H_0 : $\beta_{AUC} = \beta_{Cmax} = \log(0.8)$,

Under H_1 : $\beta_{AUC} = \beta_{Cmax} = \log(1) = 0$

- Evaluation:
 - ▶ Type I error ($IP_{95\%}(0.05) = [0.0326; 0.0729]$)
 - ▶ Power
 - ▶ Computing times

Type I error and power estimates

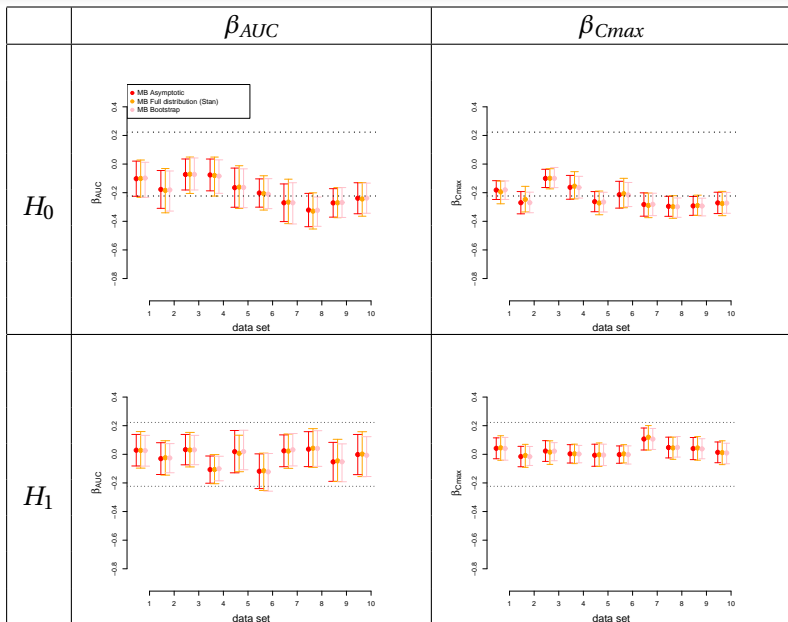
	NCA	MB Asymptotic SE	MB Bootstrap CI	MB Full distribution CI
Mean computing time for 1 data set	15 sec	15 sec	43 min	5 min

	Rich design			
	NCA	MB Asymptotic SE	MB Bootstrap CI	MB Full distribution CI
Type I error β_{AUC}	0.050	0.056	0.062	0.040
Type I error β_{Cmax}	0.062	0.058	0.064	0.044
Power β_{AUC}	0.978	0.830	0.832	0.762
Power β_{Cmax}	0.998	1.000	1.000	0.962

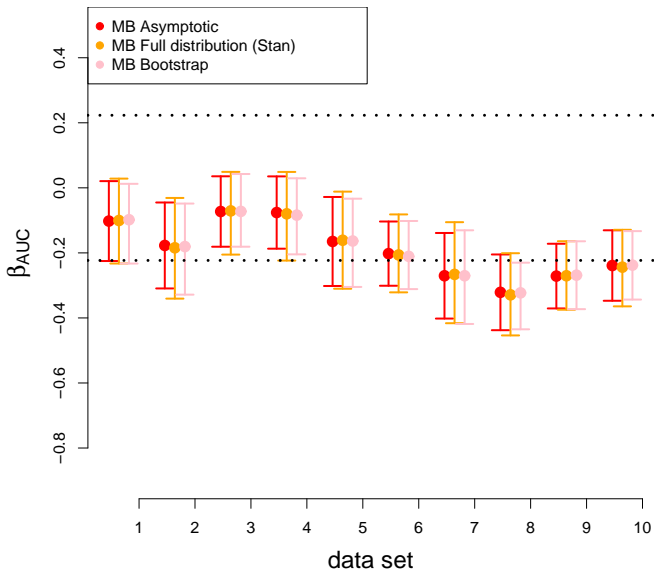
	Sparse design			
	NCA	MB Asymptotic SE	MB Bootstrap CI	MB Full distribution CI
Type I error β_{AUC}	-	0.076	0.074	0.060
Type I error β_{Cmax}	-	0.066	0.068	0.050
Power β_{AUC}	-	0.910	0.916	0.868
Power β_{Cmax}	-	1.000	1.000	0.992

$IP_{95\%}(0.05) = [0.0326; 0.0729]$ with 500 simulated data sets

90% CI(β) on sparse design



90% CI(β) on sparse design, under H0

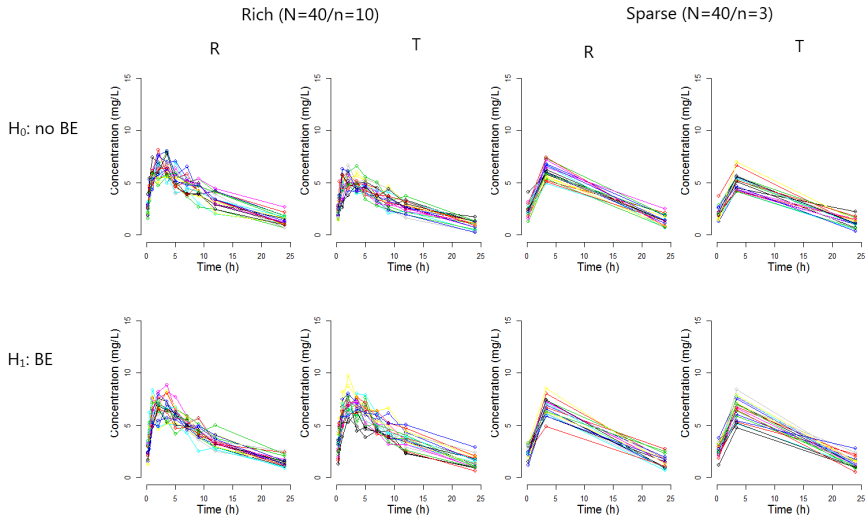


Conclusion

- Implementation of new methods for MBBE using parametric bootstrap and full distribution sampled in Stan
- Correction of MB TOST with asymptotic SE on sparse design using full distribution sampled in Stan
 - ▶ faster than bootstrap
- Perspective: implementation of the methods for 2-period 2-sequence crossover BE studies (accounting for within-subjects variability)

Thank you

Simulation scenarios (2/2)



90% CI(β) on rich design

